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Credit supply and monetary policy: Identifying the bank balance-sheet channel with loan Applications

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Bank and Firm Balance-Sheet Channels in Good and Crisis Times

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Credit Supply versus Demand:

Bank and Firm Balance-Sheet Channels in Good and Crisis Times

Abstract

Banking crises involve periods of persistently low credit and economic growth.

Banks' balance sheets are then weak but so are those of non-financial corporate

borrowers. Hence, a crucial question is whether credit growth is low due to supply or

to demand factors. However convincing identification has been elusive due to a lack

of detailed loan application-, bank-, and firm-level data. Access to a dataset of loan

applications in Spain that is matched with complete bank and firm balance-sheet data

covering the period from 2002 to 2010 allows us to identify bank and firm balance-

sheet channels. We find robust evidence showing that bank balance-sheet strength

determines the success of loan applications and the granting of loans in crisis times.

The heterogeneity in firm balance-sheet strength determines loan granting in both

good and crisis times, although the potency of this firm balance-sheet channel is the

largest in the latter period. Our findings therefore hold important implications for both

theory and policy.

Keywords: bank lending channel, credit supply, business cycle, credit crunch, capital,

liquidity.

JEL: E32, E44, E5, G21, G28.

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1. Introduction

Since 2007 Western Europe and the United States have experienced a severe

banking crisis, followed by weak credit growth and a strong economic recession.

These recent events are not unique. Banking crises are recurrent phenomena, which

trigger deep and long-lasting recessions, with depressed credit growth (Reinhart and

Rogoff (2009); Schularick and Taylor (2011)).

The main channel by which banks' balance-sheet weaknesses affect the economy

at large is through a reduction of credit supply (Bernanke (1983)). But balance sheets

of non-financial corporate borrowers may also be weak and credit demand may be

low owing to bad economic prospects (Bernanke, Gertler and Gilchrist (1996)).

Moreover, banking crises are not exogenous phenomena, but often follow a period of

strong credit growth (Kindleberger (1978); Schularick and Taylor (2011)).

A crucial question, therefore, is whether credit growth depends on supply or on

demand factors, both in good and in crisis times. Credit supply is affected by the

banks' balance-sheet strength, the so-called bank lending channel (see e.g. Bernanke

and Gertler (1987); Bernanke and Gertler (1995); Holmstrom and Tirole (1997);

Bernanke (2007); Adrian and Shin (2010); Adrian and Shin (2011); Gertler and

Kiyotaki (2011)). Demand is affected by the firm balance-sheet strength, the so-

called firm balance-sheet channel (Bernanke and Gertler (1989); Bernanke and

Gertler (1995); Bernanke, Gertler and Gilchrist (1996); Bernanke, Gertler and

Gilchrist (1999)). Bank and firm net worth vary over the business cycle, but bank net

worth and balance-sheet strength may especially matter in financial crisis times

(Gertler and Kiyotaki (2011)).

However, fully-convincing identification of the bank and firm balance-sheet

channels has remained elusive due to unavailability of detailed micro data (Bernanke

and Gertler (1995)). Spain, however, offers an ideal setting for identification: (i) As

far as we are aware, Spain is the only country where loan applications are available

for all banks and, moreover, include also an identifier for the borrower lodging the

application. Hence both bank and borrower identity are known, which is crucial to

identify credit availability; (ii) The credit application data can be matched with

comprehensive bank balance-sheet data (collected by the supervisor) and complete

firm balance-sheet data, which can proxy for the strength of bank and firm balance-

sheets. This information is essential to distinguish between the bank and the firm

balance-sheet channels.

We analyze the bank and firm balance-sheet channels using loan applications

from Spain which are matched with complete bank and firm balance-sheet data (for

example identity, size, capital ratio, liquidity ratio, credit history and defaults). The

data which is available at a monthly frequency covers the period from 2002 to 2010,

which allows us to analyze the period before the start of the financial crisis in

August 2007 which was characterized by good economic conditions (henceforth

"good times"), and the banking crisis period which started in August 2007 and we

study this period until June 2010 (henceforth "crisis times").

Analyzing first only bank balance sheet strength and loan application granting,

we find robust evidence that heterogeneity in bank balance-sheet strength does not

determine loan granting in good times. However, it does determine loan granting in

crisis times, in particular bank size, capital, liquidity, and the doubtful loan ratio. In

consequence, the estimates suggest that credit supply factors only matter in crisis

times.

In contrast, when analyzing the effect of firm balance-sheet strength on loan

granting, we find evidence that firm heterogeneity in balance-sheet strength

determines the probability a loan is granted to the applying firm both in good and in

crisis times. Firm balance-sheet strength, nevertheless, matters even more in crisis

times than in good times, with for example the impact of firm leverage on loan

application granting more than doubling in crisis times as compared to in good

times.

The key contribution of our paper is the identification strategy we employ and the

estimates we obtain. The identification of the bank versus the firm balance-sheet

channel is important both for testing theoretical models and for policy (see the last

part of the paper). There is a large empirical literature on the balance-sheet channels

that started with a macro approach and that, to achieve better identification, moved

to micro level data (Bernanke and Gertler (1995)): At the bank level to identify the

bank balance-sheet (or lending) channel (Kashyap and Stein (2000)) and at the firm

level to identify the firm balance-sheet channel (Bernanke, Gertler and Gilchrist

(1996)). The papers in the literature so far – and due to data unavailability – do not

use loan applications to analyze credit granting. In addition, as banks with different

size, net worth and risk tend to lend to firms with different size, net worth and risk

an analysis either at the bank level or at the firm level may be biased. Therefore, the

identification of the bank and firm balance-sheet channels can only be done with

loan applications matched with bank and firm identity and complete balance-sheet

¹ A large empirical literature has investigated the bank- and firm-balance sheet channels independently, with the analysis done at either the bank or the firm level. On the bank side see Bernanke and Blinder (1992), Kashyap and Stein (2000), Kishan and Opiela (2000), Jayaratne and Morgan (2000), Ashcraft (2006), Gan (2007), Khwaja and Mian (2008), Black, Hancock and Passmore (2009), and Chaney, Sraer and Thesmar (2009), among others. On the firm side, see Gertler and

Gilchrist (1994) and Bernanke, Gertler and Gilchrist (1996), for example.

data for both. As far as we are aware, this paper is the first in the literature to do so,

hereby relying on data from Spain.

This data moreover allows us, as we explain in detail in Section 2, to use

firm*time fixed effects (which control comprehensively for the firm channel) to

identify the bank balance-sheet channel;² and to use bank*time fixed effects (which

controls comprehensively for the bank channel) to identify the firm balance-sheet

channel. Importantly, our analysis suggests that not controlling adequately for either

the firm or bank channel results in biased results that are not robust to including

either firm*time or bank*time fixed effects – i.e., not controlling for the firm or bank

balance-sheet channel biases the estimates on the potency of the bank or firm

balance-sheet channel.

Another intended contribution of our paper to the literature – one we deem to be

very important – is that we analyze loan application granting both during the last

credit boom and also during the 2007-2010 crisis itself. We therefore contribute to

historical studies by Kindleberger (1978), Bernanke (1983), Reinhart and Rogoff

(2009), and Schularick and Taylor (2011) by analyzing the recent credit boom and

bust cycle with a comprehensive and unique micro data-set that allows for a better

identification of the fundamental determinants of credit during a boom and ensuing

bust.

We have used a part of the data set also in another paper. In Jiménez, Ongena,

Peydró and Saurina (2011a) we analyze the bank lending channel of monetary policy

until 2008 using the loan applications. Our innovation in this paper is two-fold: first,

² Puri, Rocholl and Steffen (2011) analyze household loan applications to German saving banks before and during the recent crisis and find that banks with exposure to US subprime assets grant less

loan applications during the crisis. However, different to our paper, they cannot control for non-financial-borrowers*time fixed effects, which exhaustively control for time-varying unobserved and

observed heterogeneity in non-financial borrower fundamentals and, as we show in this paper, are crucial not to bias the estimates of the potency of the balance-sheet channel.

crucial not to bias the estimates of the potency of the balance-sheet channel.

we analyze the (non-financial borrower) firm balance-sheet channel and compare it

to the bank balance-sheet channel, and (ii) we analyze the crisis period and compare

it to the preceding boom years. These two innovations are crucial for testing models

and for public policy analysis, and substantially differentiate our two papers.

The paper proceeds as follows. Section 2 discusses the data and the empirical

strategy. Section 3 presents and discusses the results. Section 4 concludes by

highlighting the relevant implications for theory and for public policy analysis.

2. Data and Empirical Strategy

In this Section we first discuss the data we employ in our empirical work, second

we present and discuss the empirical strategy highlighting the testable predictions

emanating from theory. Finally, we provide the definition of the dependent and

independent variables and the main econometric specification.

A. Data

We have access to The Credit Register of the Banco de España (CIR), which

contains confidential information on all business loans granted by all banks

operating in Spain (see Jiménez, Salas and Saurina (2006) for a detailed description

of the CIR).

To analyze credit demand, we focus on loan applications for commercial and

industrial (C&I) loans (82 percent of total loans) by non-financial publicly-limited

and limited-liability companies (that account for around 95 percent of all firms) to

commercial banks, savings banks and credit cooperatives (that account for more than

95 percent of the entire Spanish financial system). The dataset contains loan

applications from potential borrowers to banks that they are not currently borrowing

³ Delgado, Salas and Saurina (2007) explain the main features of the Spanish banking system.

from (i.e., the extensive margin of new lending). Loan applications are available

since 2002:01. Though the applications can be made at any time, they are collated

monthly and uniquely link borrowers with banks (Jiménez et al. (2011a) provides a

detailed description of this dataset).

We analyse the loan applications until 2010:06, the time of the start of the

sovereign debt crisis in the Euro area. For each loan application between 2002:02

and 2010:06, we also observe whether the loan is accepted and granted, or not, by

matching the loan application database with the CIR database, which contains the

stock of all the loans granted. Therefore, if multiple banks request information on a

particular borrower, we can infer the bank that granted the loan and the banks that

did not. In case there is a loan application but the bank does not grant the loan, either

the bank denied the firm credit or the firm perceived the offered conditions by the

bank to be less attractive than those of the loan it eventually took. Hence, we can

link loan granting for the same firm within each month to bank balance-sheet

strength, and we can also analyse the success of a loan application depending on the

firm balance-sheet strength.

We therefore match the application dataset with bank and firm datasets, so that

we have balance-sheet information for each bank that receives a loan application and

for the firm that applies for a loan. The banks' dataset, at a monthly frequency

starting in 1984, is owned by the *Banco de España* in its role as banking supervisor.

The firms' dataset is available from the Spanish Mercantile Register at a yearly

frequency, starts in 1992 and covers the large majority of firms. We can match

427,364 loan applications to bank balance-sheet data and 198,350 loan applications

to both bank and firm balance-sheet data, which constitute our two samples of loan

applications that we analyze.4

B. Empirical Strategy

The theory of the bank and firm balance-sheet channels that we discuss in the

Introduction has the following testable predictions. Bank and firm variables that

proxy for the strength of balance sheets determine loan application granting, and the

impact is stronger in crisis than in good times. Given that the main problem in the

literature is to identify the channels, we emphasize more the empirical strategy and

the data that is needed to test the predictions emanating from the theoretical

literature.

As we have access to loan applications plus the bank and firm balance-sheet

characteristics that determine balance-sheet strength, we are able to better

disentangle the supply from the demand for loans. Through the loan applications,

loan demand for each bank is in a sense given and observed, and each bank has to

decide only on the granting of each loan knowing the firm. As far as we are aware,

ours is the first paper that analyzes the impact of the bank versus the firm balance-

sheet channel relying on the probability loans are granted following applications

from firms.

To analyze the bank and firm balance-sheet channels we exploit the cross-

sectional implications of the sensitivity of credit availability in good and crisis times

according to the strength of the balance sheets (see e.g. Kashyap and Stein (2000) for

⁴ In case there is no balance-sheet information for a firm, there is nevertheless the firm identity, which is crucial to identify credit supply as we discuss below. The loan applications which are not matched to the firm balance-sheet data are from very small firms since CIR collects all business loans

from all the firms in Spain, including the very small ones.

the bank balance-sheet (or lending) channel and Bernanke, Gertler and Gilchrist

(1996) for the firm balance-sheet channel).

Following the theoretical literature (Holmstrom and Tirole (1997), Bernanke,

Gertler and Gilchrist (1999), and Gertler and Kiyotaki (2011)) we focus on bank and

firm capital ratios.⁵ Since risk also affects net worth, we also feature for banks a non-

performing ("doubtful") loan ratio and a Herfindahl-Hirschman index of the bank's

credit portfolio by industry that proxies for bank diversification. For firms we feature

a measure of previous bad credit history and the age of the firm. Following Kashyap

and Stein (2000) and Bernanke, Gertler and Gilchrist (1996) we also feature the bank

and firm liquidity ratios and size.

We control with bank and firm fixed effects for time-invariant characteristics of

banks and firms respectively, such as for example bank type (i.e., commercial,

savings or cooperative) and firm legal structure, industry and location. Moreover,

given that banks of different net worth may be approached by borrowers with

different net worth and risk, our benchmark regressions have the largest set of

possible controls: Firm*month fixed effects to identify the bank balance-sheet

channel and bank*month fixed effects to identify the firm balance-sheet channel.

Firm*month fixed effects are a complete set of monthly dummies (from 2002:02

to 2010:06) for each firm, which therefore control exhaustively for all time-varying

observed and unobserved firm heterogeneity. This set of effects is key to control for

the demand side, and hence, to identify the bank balance-sheet channel. Bank*month

fixed effects are a complete set of monthly dummies (from 2002:02 to 2010:06) for

⁵ Off-balance sheet volumes are very small in Spain. Hence, total bank assets cover most of the banks' businesses. Banks did not develop conduits and/or Structured Investment Vehicles (SIVs) because the prevailing accounting rules made banks consolidate these vehicles and set aside sufficient

capital, eliminating the incentives of banks for developing such structures.

each bank, which therefore control exhaustively for all time-varying observed and

unobserved bank heterogeneity. Similarly this set of effects is key to control for the

supply side, and hence, to identify the firm balance-sheet channel.

Given that these comprehensive sets of fixed effects does not allow including

concurrently respectively both firm or bank balance sheet variables, we also analyze

specifications without these sets of fixed effects. In these cases, i.e., when there are

no time fixed effects included, we control for macro factors with real GDP growth,

the change in the interbank 3-month interest rate, and inflation (using the Consumer

Price Index).

C. Dependent Variable, Independent Variables and Specifications

In this subsection we provide the definition of the main dependent variable, the

independent variables, and the estimated specifications.

1. Main Dependent Variable: LOAN APPLICATION IS GRANTED

Table 1 defines the dependent and independent variables employed in the first set

of empirical specifications where we only analyze the bank balance-sheet channel,

and Table 4 defines the second set, where we analyze both the bank and the firm

balance-sheet channels. Tables 1 and 4 also present their descriptive statistics for the

whole period (2002:02-2010:06), for the good times (2002:02-2007:07), and for the

crisis times (2007:08-2010:06).

The dependent variable we feature throughout the paper is LOAN

APPLICATION IS GRANTED (we recurrently shorthand this as "loan granting"),

which equals one if the loan application by firm i at time t is approved by bank b and

the loan is granted in month t to t+3, and equals zero otherwise. The average value

of loan granting equals 39 percent in both Tables 1 and 4 in good times and 30

percent in crisis times.⁶

We match each loan application with its relevant bank and firm characteristics, in

particular firm identity. The inclusion of firm (or firm-month) fixed effects in a logit

(or probit) model naturally restricts the sample to those firms that filed at least one

application that did result in a loan granted and one application that did not during

the sample period (or in a month). To avoid this selection problem we employ linear

probability models in the regressions.

2. Independent Variables

As independent variables we include an array of bank and firm characteristics that

proxy for bank and firm balance-sheet strength. The summary statistics of Table 1

are based on the observations used in the first three Tables that include only bank

characteristics (and possibly firm fixed effects or firm*time fixed effects). Bank

balance-sheet data is taken at the end of the previous month t-1.

The bank balance-sheet variables we are foremost interested in are bank size,

capital, liquidity and risk. Bank size is the log of the total assets of the bank, BANK

LN(TOTAL ASSETS), its average is 17.27 in good times and 17.71 in crisis times

(31 and 49 billion Euros, respectively). The BANK CAPITAL RATIO as a measure

of the bank's net worth which is defined as the ratio of core capital over total assets

of the bank (as in Bernanke and Lown (1991) for example). Core capital is defined

as total equity plus retained earnings. As we use the book value of equity and assets

are not risk adjusted, our measure is equivalent to a pure leverage ratio. Thus defined

it has an average value of 5.47 percent in good times and 5.39 percent in crisis times.

⁶ As we explained above, there is no firm balance-sheet data for some loan applications but firm identity is always known and so is complete bank balance-sheet data.

We also use a measure of banks' liquidity position. The BANK LIQUIDITY

RATIO is the ratio of liquid assets held by the bank (i.e., cash and deposits with

central banks and other credit institutions, and public debt with a maturity up to one

year) and the total assets of the bank. Banks on average held 17.14 percent of their

balance-sheet in liquid assets in good times but only 12.51 percent in crisis times.

We proxy bank risk by the doubtful loan ratio of the BANK DOUBTFUL

LOANS RATIO which has an average value of 0.73 percent in good times and 2.71

in crisis times, and by the BANK HERFINDAHL BY INDUSTRY, the Herfindahl-

Hirschman index of the bank's credit portfolio by industry, which has an average

value of 27.26 in good times and 28.58 in crisis times.

As a bank-firm relationship variable we include LN(1+NUMBER OF MONTHS

WITH THE BANK), which is the log of one plus the number of months that the

bank had a working relationship with the firm.

To analyze the firm balance-sheet channel we include a broad set of firm

characteristics that proxy for the strength of firm balance sheets (see Table 4 for the

summary statistics). Parallel to the bank variables, as firm variables we feature:

FIRM LN(TOTAL ASSETS), the log of the total assets, which has a value of 7.65 in

good times and 7.74 in crisis times (2 and 2.2 million Euros, respectively); FIRM

CAPITAL RATIO, which is the log of the ratio of own funds over total assets of the

firm, which has an average value of 2.57 in good times and 2.84 in crisis times; the

FIRM LIQUIDITY RATIO, the current assets over total assets of the firm, which

has an average value of 6.71 in good times and 7.09 in crisis times.

For firm risk we use FIRM SUBPRIME, a dummy variable that equals one if the

firm had delinquent loans before the loan was requested, and equals zero otherwise.

Its average value equals 10 percent in good times and 12 percent in crisis times; and

FIRM LN(1+AGE), the log of one plus the age of the firm in years that has an

average value of 1.96 in good times and 2.26 in crisis times (7 and 9 years,

respectively).

3. Specifications

The specifications we estimate are at the loan application-level and we match the

loan application outcomes (whether the loan is granted or not) with the associated

bank and firm balance-sheet variables. We analyze first good times (2002:02-

2007:07) and then we analyze the whole period (2002:02-2010:06) introducing a

crisis dummy variable that takes the value of one after 2007:07 and its interactions

with the bank and firm balance sheet variables. We provide in the next Section the

exact empirical specification we discuss in each Table, but the most general

empirical specification assessing the probability a loan application is granted is

structured as follows:

LOAN APPLICATION IS GRANTED_{bit} =

 $bank_{bt-1} + firm_{it-1} + CRISIS_{t-1} + CRISIS_{t-1} * bank_{bt-1} + CRISIS_{t-1} * firm_{it-1} + fixed$ (1)

effects + $\varepsilon_{\rm bit}$,

where bank and firm are respectively the bank and the firm balance-sheet

variables presented above, CRISIS is the crisis dummy that takes the value of one in

the sample months after 2007:07 and equals zero otherwise, CRISIS * bank and

CRISIS * firm are the interactions between the dummy crisis and the firm and bank

balance-sheet variables, fixed effects are the different specifications of bank, time

and firm fixed effects we presented above, in particular bank, time, firm, firm*time,

and bank*time fixed effects. The theory of the bank and firm balance-sheet channels

predict that bank and firm variables proxying for balance-sheet strength matter, and

especially in crisis times (i.e., when CRISIS = 1).

3. Results

We first analyze the bank balance-sheet channel with the sample composed by all

loan applications (see Tables 1 to 3), and then we analyze the bank and the firm

balance-sheet channels with the sample of loan applications matched to both bank

and firm balance-sheet data (see Tables 4 to 6).

A. The Bank Balance-Sheet Channel

Table 1, as explained in the previous Section, provides in addition to the variable

definitions also the summary statistics. Table 2 provides the results for the bank

balance-sheet channel for the period of good times and Table 3 for the whole period.

The specifications we estimate are as follows:

LOAN APPLICATION IS GRANTED_{bit} =

(2)

 $bank_{bt-1} + controls_{bit} + fixed \ effects + \varepsilon_{bit}$

where bank includes BANK LN(TOTAL ASSETS), BANK CAPITAL RATIO,

BANK LIQUIDITY RATIO, BANK DOUBTFUL LOANS RATIO, and BANK

HERFINDAHL BY INDUSTRY; and controls include LN(1+NUMBER OF

MONTHS WITH THE BANK), GDP GROWTH, CHANGE IN 3-MONTH

INTEREST RATE, and INFLATION. The latter three variables drop out when we

include time effects starting in Model 2.

In Table 2, when we do not control for time or firm fixed effects yet, but only

include bank fixed effects (Model 1), we find that smaller banks grant loans with a

higher probability than larger banks. The estimated coefficient equals -3.61***.

Given that we estimate linear probability models and given that the estimated

coefficients are expressed in percent, the economic magnitude of the effect can be

readily approximated. A decrease in bank asset size of one standard deviation (i.e.,

1.47), increases the probability a loan application is granted by 5.2 percentage points

(= 1.45 times 3.61***). This is a sizeable effect given that the probability that a loan

application is accepted in good times equals 39 percent, implying a semi-elasticity of

13 percent.

Banks with a more diverse loan portfolio also grant loans with a higher

probability, but the economic magnitude of the effect is somewhat smaller: A one

standard deviation lower concentration results in a 1.3 percentage points increase

(9.40 times -0.14**) in granting loans. This finding implies that in good times banks

that are diversified across industries are more likely to grant loans.

Finally, we find that banks are more likely to grant loans to firms with a longer

previous relationship (but the economic relevancy is also modest) and that a one

percentage point higher GDP growth implies a 2.9 percentage points higher

probability of loan granting (= 1 times 2.92***).

Model 2 adds time fixed effects. The estimated coefficient on bank size turns

statistically insignificant. In Model 3 we add firm fixed effects and bank liquidity

⁷ As in the Tables, ***, **, and * indicates statistical significant at the 1, 5, and 10 percent level,

respectively.

becomes marginally statistically significant. Time and firm fixed effects control

partially for loan demand net worth and risk and, as the change of results suggest,

they are necessary as for example banks with different size and liquidity likely have

different type of borrowers.

Yet to fully control for credit demand we need to control for both time-varying

observed and unobserved heterogeneity in firm balance-sheet strength (since the

business and monetary cycle affect credit demand). We do this in Model 4 where we

add firm*time fixed effects in addition to bank fixed effects. Now only the estimated

coefficients on bank concentration and the bank-firm relationship variables are now

statistically significant, but as calculated earlier in Model 1 their economic relevancy

is rather modest.

In sum, analyzing the effect of bank balance sheet strength on the probability a

loan is granted following applications, the evidence suggests that the heterogeneity

in bank balance-sheet strength (i.e., bank size, capital, liquidity, and risk) does not

determine loan granting in good times.

In Table 3 we use loan applications from the whole period and through

interactions of the crisis dummy with bank balance-sheet variables aim to

differentiate the impact of bank balance-sheet strength on lending in normal versus

crisis times. As the crisis shock was unexpected, it is difficult to believe that banks

already adjusted their balance sheets anticipating the crisis.8

The specifications we now estimate are:

.

⁸ Our results are similar if during the crisis period we use the relevant values for bank and firm characteristics immediately prior to 2007:08.

LOAN APPLICATION IS GRANTED_{bit} =

(3)

 $bank_{bt-1} + CRISIS_{t-1} + CRISIS_{t-1} * bank_{bt-1} + controls_{bit} + fixed\ effects + \varepsilon_{bit}$

where bank and controls include the same set of variables as in (2).

The crisis drastically decreases the probability a loan application is successful.

Likely concurrent lower GDP growth, a higher short-term interest rate, and higher

inflation also result in a lower probability of loan granting.

When analyzing the "strongest" specification, our benchmark Model 4, which is

saturated with comprehensive sets of bank and firm*time fixed effects, we find that

bank balance-sheet variables do not matter in normal times, but do matter in crisis

times. In particular, banks that are smaller, with lower capital ratios, or with more

doubtful loans are less likely to grant loans in crisis times. Banks also tend to grant

more loan applications to firms which they had lent in the past, but this effect is not

different between crisis and good times.

The economic relevancy of the estimated effects of the bank balance sheet

strength is sizable. For example for a one standard deviation commensurate change

in crisis times in bank size (decrease), capital (decrease), or doubtful loan ratio

(increase) the probability a bank loan application is granted decreases by 1.0, 1.8 and

2.3 percentage points, respectively (-1.46 times 0.66***; -1.84 times 0.95***; 2.27

times -1.03**).9 As the probability a loan application is granted in crisis times equals

30 percent, the semi-elasticities amount to 3.3, 6 and 7 percent, respectively.

⁹ Rochet and Vives (2004) and Vives (2011b) show that low bank net worth (capital and doubtful ratio) negatively affect bank liquidity, especially during crisis times (leading, therefore, to a reduction in bank assets, in particular new credit). See also Gale and Yorulmazer (2011).

In sum, analyzing multiple loan applications from the same borrower in the same

month (firm*time fixed effects), and accounting for all time-invariant bank

characteristics, banks with stronger balance-sheets grant loan applications more

readily than banks with weaker balance-sheets in crisis times, but not in good times.

Hence the results suggest that credit supply factors only matter in crisis times. Not

controlling exhaustively for the firm balance-sheet channel biases the estimates of

the potency of the bank lending channel.

B. The Bank and Firm Balance-Sheet Channels

Table 4, as explained in Section 2, provides the summary statistics for the loan

applications that are also matched with firm balance sheets. Table 5 provides the

results for the period of good times for the bank and firm balance-sheet channels and

Table 6 for the whole period.

The specifications we now estimate take the form:

LOAN APPLICATION IS GRANTED_{bit} =

(4)

 $bank_{bt-1} + firm_{it-1} + controls_{bit} + fixed \ effects + \varepsilon_{bit}$

where bank and controls include the same set of variables as in (2) and (3), while

firm includes: FIRM LN(TOTAL ASSETS), FIRM CAPITAL RATIO, FIRM

LIQUIDITY RATIO, FIRM SUBPRIME, and FIRM LN(1+AGE).

In Tables 5 and 6 we follow the structure of the previous Tables and

progressively saturate the specification with comprehensive sets of fixed effects, i.e.,

we introduce comprehensive sets of bank, time, firm and/o bank*time effects.

In Models 1 and 2, without controlling for firm fixed effects, we find similar

results for the bank variables as in Table 2 and we also find that smaller firms, with a

lower capital ratio or that are younger have higher probability of being successful in

their loan application.

In Model 3 we control for firm fixed effects (in addition to bank and time fixed

effects that were introduced in Models 1 and 2, respectively). We now find that firms

with a higher capital ratio and with a better credit history have higher loan granting

probability, and we still find that smaller firms obtain higher credit granting. We also

find similar results for bank variables as in Model 3 of Table 2.

In Models 4 and 5 we introduce bank*time fixed effects in addition to the firm

fixed effects to control for time-varying observed and unobserved heterogeneity in

bank balance-sheet strength. Given the large set of fixed effects we cannot double

cluster: Model 4 therefore provides the results with firm clustering and Model 5 with

bank clustering. We find that in good times firms with higher capital ratio and with a

better credit history have a higher probability their application will be resulting in a

loan being granted.

The estimated effects are also economically relevant. A one standard deviation

increase in the firm capital ratio results in a 3 percentage points increase in the

probability (1.16 times 2.56***), a semi-elasticity of 8 percent (3 divided by 39).

Firms that are prime have a 7 percentage points higher probability of getting a loan

upon applying than subprime firms.

Importantly as well, as in Tables 2 and 3 for the bank channel, the results imply

that not controlling for firm fixed effects or bank*time fixed effects biases the

estimates of the potency of the firm balance-sheet channel, in particular the elasticity

of firm capital and subprime without firm fixed effects, and of firm size without

bank*time fixed effects.

Next in Table 6 we estimate specifications of the form:

LOAN APPLICATION IS GRANTED_{bit} =

 $bank_{bt-1} + firm_{it-1} + CRISIS_{t-1} + CRISIS_{t-1} * bank_{bt-1} + CRISIS_{t-1} * firm_{it-1} + (5)$

 $controls_{bit} + fixed\ effects + \varepsilon_{bit}$,

where bank, firm and controls include the same set of variables as were

introduced in (2) and (3), and (5), respectively.

In the benchmark regressions including firm fixed effects in conjunction with

bank*time fixed effects (columns 4 and 5), we find that firms with a lower capital

ratio are less likely to obtain credit in general, but that the effects are stronger in

crisis times. A one standard deviation decrease in the firm capital ratio in good times

lowers the probability by 2 percentage points (-1.11 times 1.51***), and a similar

decrease in the capital ratio in crisis times lowers the probability by an additional 3

percentage points (-1.16 times 2.40***), implying a total semi-elasticity of 17

percent (5 divided by 30).

Younger firms are also less likely to obtain credit following an application in the

crisis times with a similarly sized economic effect. Interestingly, subprime firms are

penalized equally in crisis versus good times. Finally, and similarly as in Table 3 but

differently to the firm balance-sheet channel, the bank balance-sheet strength does

not matter in good times, but it does in crisis times (see Models 1 to 3).¹⁰

In sum, results suggest that heterogeneity in bank balance-sheet strength does not

determine loan granting in good times. However, it does determine loan granting in

the crisis. In consequence, the results suggest that credit supply factors only matter

in crisis times.

When analyzing firm balance-sheet strength, we instead find evidence that firm

heterogeneity in balance-sheet strength determines loan application granting both in

good and crisis times. Firm balance-sheet strength, nevertheless, matters more in

crisis than in good times, with key elasticities as firm leverage to loan application

granting more than doubling in crisis times as compared to good times.

Finally, the results imply unequivocally that not controlling exhaustively for the

firm balance-sheet channel biases the estimates of the potency of the bank balance-

sheet channel, and that similarly not controlling exhaustively for the bank balance-

sheet channel biases the estimates of the potency of the firm balance-sheet channel.

4. Conclusions and Implications for Theory and Policy

The recent crisis has resulted in massive transfers from governments and central

banks to banks, through government bail-outs, recapitalizations and liquidity

assistance and various central bank lender-of-last-resort actions to help banks in

repairing their capital and liquidity positions. Our evidence shows that weaknesses in

bank balance sheets reduces the supply of bank credit in crisis times (credit crunch)

and, therefore, our estimates lend support to theories that emphasize the role of

¹⁰ Notice that we do not control for firm*time fixed effects in Table 6 and that the coefficient on the bank doubtful loan ratio is not significant, see Table 3 Model 3 versus 4 where this coefficient was only statistically significant when we introduce firm*time fixed effects and the coefficient increases

from to Model 3 to 4 by a factor of almost ten.

banks for the business cycle and crises (see e.g. Holmstrom and Tirole (1997); Allen

and Gale (2007); Matsuyama (2007); Shleifer and Vishny (2010b); Shleifer and

Vishny (2010a); Adrian and Shin (2011); Gertler and Kiyotaki (2011); Diamond and

Rajan (2011), Vives (2011b)).

Firm balance-sheet strength matters in general but effects are stronger in crisis

times as highlighted by our estimates of the potency of the firm-balance sheet

channel (Bernanke, Gertler and Gilchrist (1996); Bernanke, Gertler and Gilchrist

(1999); Kiyotaki and Moore (1997); Lorenzoni (2008); Jeanne and Korinek (2010)).

A crucial firm balance sheet characteristic that matters in the crisis is firm leverage,

in particular high leverage, which lends support to the theories of firm debt overhang

and deleveraging (see e.g. Myers (1977)). This implies that even if the government

support to banks helps bank credit availability, firms' balance-sheet strength and

access to finance is also important. Therefore, our results support some of the

policies by the Federal Reserve targeted to non-financial borrowers' access to

finance.

Though our results indicate that heterogeneity in bank balance-sheet do not

determine loan application granting in good times, it does not mean that banks are

irrelevant for credit built-up in good times. Risk-taking incentives captured by

changes in *composition* in credit supply could be more important in good times

(Jiménez, Ongena, Peydró and Saurina (2011b), Allen and Rogoff (2011)). Finally,

our findings that bank strength does significantly matter in crisis times for lending

policies supports current work by regulators to strengthen capital and liquidity levels

at each individual bank, so that when the next crisis arrives banks are in a better

position to cope with it and, thus, the crisis will have a attenuated impact on credit granting (i.e., a "softer" credit crunch) and, therefore, on the real economy.¹¹

¹¹ Support in the literature for stronger regulatory requirements can be found in Admati, DeMarzo, Hellwig and Pleiderer (2010), Hellwig (2010), Repullo, Saurina and Trucharte (2010) and Hanson, Kashyap and Stein (2011), among others. Insurance contracts contingent on average bank capital as a way to insure against systemic crises are discussed in Gersbach (2011). For a discussion on competition and the limits to regulation see Vives (2011a).

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Table 1

Table 1					
Variable definitions and descriptive statistics for all loan	applications				
Variable Names	Variable Definition				
LOAN APPLICATION IS GRANTED	A dummy variable which equals one if the loan application by firm i at time (i.e., month) t is approved by bank b and the loan is granted in month t to $t+3$, and equals zero otherwise				
BANK LN(TOTAL ASSETS)	=		n thousands of Eur	ros	
BANK CAPITAL RATIO	The log of total assets of the bank, in thousands of Euros The ratio of bank equity over total assets of the bank, in percent				
BANK LIQUIDITY RATIO	The ratio of liquid				nd loans and
Britik Eligoidi Fikitio	advances to govern assets of the bank, ir	ments and credi			
BANK DOUBTFUL LOANS RATIO	The doubtful loan ra	tio of the bank,	in percent		
BANK HERFINDAHL BY INDUSTRY	The Herfindahl-Hirs	schman index of	the bank's credit	portfolio by ind	ustry
LN(1+NUMBER OF MONTHS WITH THE BANK)	The log of one plus months				
CRISIS	A dummy variable otherwise	which equals of	one in months aft	ter 2007:07 and	d equals zero
GDP GROWTH	Annual change of Sp	oanish gross dor	nestic product in r	eal terms, in pe	rcent
CHANGE IN 3-MONTH INTEREST RATE	Annual change of Sp		-	-	
INFLATION	Annual change of S ₁				
Descriptive Statistic		St. Dev.	Minimum	Median	Maximum
Whole Period (2002:02-2010:06)		servations = 42	7,364		
LOAN APPLICATION IS GRANTED	0.35	0.48	0	0	1
BANK LN(TOTAL ASSETS)	17.49	1.47	9.60	17.66	19.94
BANK CAPITAL RATIO	5.43	2.03	0.00	4.91	87.17
BANK LIQUIDITY RATIO	14.90	7.53	0.03	14.00	92.07
BANK DOUBTFUL LOANS RATIO	1.69	1.93	0	0.82	58.61
BANK HERFINDAHL BY INDUSTRY	27.90	8.84	13.20	25.62	100
LN(1+NUMBER OF MONTHS WITH THE BANK)	0.38	1.11	0	0	5.67
CRISIS	0.48	0.50	0	0	1
GDP GROWTH	1.71	2.71	-4.45	3.02	4.11
CHANGE IN 3-MONTH INTEREST RATE	-0.28	1.53	-4.38	0.09	1.41
INFLATION	2.68	1.56	-1.37	2.92	5.27
Good Times (2002:02-2007:07)	Number of Ob	servations = 22	0,275		
LOAN APPLICATION IS GRANTED	0.39	0.49	0	0	1
BANK LN(TOTAL ASSETS)	17.27	1.45	9.60	17.40	19.71
BANK CAPITAL RATIO	5.47	2.20	0.00	4.91	63.10
BANK LIQUIDITY RATIO	17.14	8.10	0.03	15.40	92.07
BANK DOUBTFUL LOANS RATIO	0.73	0.66	0	0.54	28.33
BANK HERFINDAHL BY INDUSTRY	27.26	9.40	13.20	24.01	87.94
LN(1+NUMBER OF MONTHS WITH THE BANK)	0.32	1.01	0	0	5.56
GDP GROWTH	3.56	0.42	2.57	3.65	4.11
CHANGE IN 3-MONTH INTEREST RATE	0.28	0.81	-1.40	0.19	1.31
INFLATION	3.11	0.62	2.13	3.15	4.19
Crisis Times (2007:08-2010:06)		servations = 20			
LOAN APPLICATION IS GRANTED	0.30	0.46	0	0	1
BANK LN(TOTAL ASSETS)	17.71	1.46	9.94	17.87	19.94
BANK CAPITAL RATIO	5.39	1.84	0.00	4.91	87.17
BANK LIQUIDITY RATIO	12.51	6.02	0.53	11.74	90.76
BANK DOUBTFUL LOANS RATIO	2.71	2.27	0	2.23	58.61
BANK HERFINDAHL BY INDUSTRY	28.58	8.15	13.65	26.89	100
LN(1+NUMBER OF MONTHS WITH THE BANK)	0.44	1.21	0	0	5.67
GDP GROWTH	-0.27	2.73	-4.45	-0.49	3.58

-0.87

2.22

1.86

2.05

-4.38

-1.37

-0.50

1.77

1.41

5.27

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CHANGE IN 3-MONTH INTEREST RATE

INFLATION

Table 2
The bank balance-sheet channel during good times (2002:02-2007:07) using all loan applications

	(1)	(2)	(3)	(4)
BANK LN(TOTAL ASSETS)	-3.61 ***	-1.74	-0.60	-0.27
	(1.38)	(1.76)	(1.85)	(1.74)
BANK CAPITAL RATIO	-0.10	-0.06	0.07	0.30
	(0.30)	(0.33)	(0.27)	(0.27)
BANK LIQUIDITY RATIO	-0.11	-0.15	-0.17 *	-0.13
	(0.13)	(0.13)	(0.10)	(0.08)
BANK DOUBTFUL LOANS RATIO	-0.22	-0.20	0.18	0.24
	(0.68)	(0.68)	(0.61)	(0.54)
BANK HERFINDAHL BY INDUSTRY	-0.14 **	-0.13 *	-0.14 **	-0.15 ***
	(0.07)	(0.08)	(0.06)	(0.06)
LN(1+NUMBER OF MONTHS WITH THE BANK)	1.21 ***	1.22 ***	0.99 ***	1.10 ***
	(0.16)	(0.16)	(0.20)	(0.19)
GDP GROWTH	2.92 ***			
	(1.05)			
CHANGE IN 3-MONTH INTEREST RATE	-0.13			
	(0.55)			
INFLATION	0.25			
	(0.21)			
Fixed Effects				
Ban	k Yes	Yes	Yes	Yes
Tim	e No	Yes	Yes	-
Fire	n No	No	Yes	-
Firm*Tim	e No	No	No	Yes
Number of Observations	220,275	220,275	220,275	220,275

Notes: The table reports the estimated coefficients and robust standard errors (S.E.) in percent clustered at the bank and firm level from linear probability models estimated using least squares. The dependent variable is LOAN APPLICATION IS GRANTED. Fixed effects are included ("Yes"), not included ("No"), or comprised by the included set of fixed effects ("-"). The set of time fixed effects includes a fixed effect for every (but one) year:month during the sample period. The variable definitions and summary statistics are in Table 1. ***, **, and * indicates statistical significant at the 1, 5, and 10 percent level, respectively.

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Table 3
The bank balance-sheet channel during good (2002:02-2007:07) and crisis times (2007:08-2010:06) using all loan applications

	(1)	(2)	(3)	(4)
BANK LN(TOTAL ASSETS)	0.23	0.04	-0.11	-0.10
	(2.20)	(3.49)	(2.85)	(2.19)
BANK CAPITAL RATIO	-0.30	-0.44	0.15	-0.02
	(0.38)	(0.41)	(0.34)	(0.31)
BANK LIQUIDITY RATIO	-0.08	-0.08	-0.02	-0.04
	(0.11)	(0.11)	(0.08)	(0.07)
BANK DOUBTFUL LOANS RATIO	0.26	0.15	-0.73 ***	0.35
	(0.72)	(0.73)	(0.22)	(0.46)
BANK HERFINDAHL BY INDUSTRY	-0.09	-0.11	-0.08	-0.09
	(0.08)	(0.09)	(0.07)	(0.05)
LN(1+NUMBER OF MONTHS WITH THE BANK)	1.22 ***	1.23 ***	0.86 ***	1.11 ***
	(0.18)	(0.17)	(0.15)	(0.19)
CRISIS	-22.23 ***	, ,	, ,	, ,
	(7.64)			
CRISIS * BANK LN(TOTAL ASSETS)	0.75 **	0.70 **	0.01	0.66 ***
,	(0.32)	(0.31)	(0.07)	(0.24)
CRISIS * BANK CAPITAL RATIO	0.99 ***	0.95 ***	0.32 ***	0.95 ***
	(0.24)	(0.26)	(0.10)	(0.17)
CRISIS * BANK LIQUIDITY RATIO	0.13	0.10	0.05	0.06
· ·	(0.12)	(0.12)	(0.06)	(0.07)
CRISIS * BANK DOUBTFUL LOANS RATIO	-0.60	-0.76	-0.12	-1.03 **
	(0.75)	(0.78)	(0.12)	(0.47)
CRISIS * BANK HERFINDAHL BY INDUSTRY	-0.07	-0.07	-0.08 **	-0.07
	(0.08)	(0.08)	(0.03)	(0.05)
CRISIS * LN(1+NUMBER OF MONTHS WITH THE BANK)	0.14	0.13	0.07	-0.16
,	(0.16)	(0.16)	(0.13)	(0.22)
GDP GROWTH	1.84 ***	, ,	, ,	` ,
	(0.37)			
CHANGE IN 3-MONTH INTEREST RATE	-1.18 **			
	(0.51)			
INFLATION	-0.46 **			
	(0.20)			
Fixed Effects	, ,			
Bank	Yes	Yes	Yes	Yes
Time	No	Yes	Yes	-
Firm	No	No	Yes	-
Firm*Time	No	No	No	Yes
Number of Observations	427,364	427,364	427,364	427,179

Notes: The table reports the estimated coefficients and robust standard errors (S.E.) in percent clustered at the bank and firm level from linear probability models estimated using least squares. The dependent variable is LOAN APPLICATION IS GRANTED. Fixed effects are included ("Yes"), not included ("No"), or comprised by the included set of fixed effects ("-"). The set of time fixed effects includes a fixed effect for every (but one) year:month during the sample period. The variable definitions and summary statistics are in Table 1. ***, **, and * indicates statistical significant at the 1, 5, and 10 percent level, respectively.

7	Γable 4					
7	Variable definitions and descriptive statistics for the loan a	applications that are a	also matched wit	th firm balance she	eet data	
F	Firm Variable Names	Variable Definition				
F	FIRM LN(TOTAL ASSETS)	The log of total asse	ets of the firm, ir	n thousands of Eur	os	
F	FIRM CAPITAL RATIO	The log ratio of firm	n own funds ove	r total assets of th	e firm	
F	FIRM LIQUIDITY RATIO	The ratio of current				
г	FIRM SUBPRIME	A dummy variable	which aquals o	na if tha firm was	delinguent on	a loan before
1	TRIVI SUDI RIIVIE	and equals zero other	•	ne ii the iiiii was	s definquent on	a loan before
	EIDM LN/1 - A CIEV	1		a car		
1	FIRM LN(1+AGE)	The log of one plus				34 :
_	Descriptive Statistics		St. Dev.	Minimum	Median	Maximum
_	Whole Period (2002:02-2010:06)		oservations = 19	•	0	
	LOAN APPLICATION IS GRANTED	0.35	0.48	0	-	10.04
	BANK LN(TOTAL ASSETS)	17.49	1.47	9.60	17.66	19.94
	BANK CAPITAL RATIO	5.43	2.03	0.00	4.91	87.17
	BANK LIQUIDITY RATIO	14.90	7.53	0.03	14.00	92.07
	BANK DOUBTFUL LOANS RATIO	1.69	1.93	0	0.82	58.61
	BANK HERFINDAHL BY INDUSTRY	27.90	8.84	13.20	25.62	100
	LN(1+NUMBER OF MONTHS WITH THE BANK)	0.38	1.11	0	0	5.67
	FIRM LN(TOTAL ASSETS)	7.70	1.68	1.45	7.65	15.82
	FIRM CAPITAL RATIO	2.70	1.11	-5.12	2.88	4.61
	FIRM LIQUIDITY RATIO	6.90	11.01	0	2.81	100
	FIRM SUBPRIME	0.09	0.29	0	0	1
	FIRM LN(1+AGE)	2.11	0.91	0	2.20	4.92
	CRISIS	0.48	0.50	0	0	1
	GDP GROWTH	1.71	2.71	-4.45	3.02	4.11
	CHANGE IN 3-MONTH INTEREST RATE	-0.28	1.53	-4.38	0.09	1.41
	NFLATION	2.68	1.56	-1.37	2.92	5.27
_	Good Times (2002:02-2007:07)		oservations = 10			
	LOAN APPLICATION IS GRANTED	0.39	0.49	0	0	1 10.71
	BANK LN(TOTAL ASSETS)	17.26	1.42	9.60	17.39	19.71
	BANK CAPITAL RATIO	5.44	2.28	0.00	4.90	63.10
	BANK LIQUIDITY RATIO	17.68	8.30	0.03	15.82	92.07
	BANK DOUBTFUL LOANS RATIO	0.72	0.64	0	0.54	28.33
	BANK HERFINDAHL BY INDUSTRY	26.92	9.34	13.74	23.43	87.94
	LN(1+NUMBER OF MONTHS WITH THE BANK)	0.46	1.19	0	0	5.56
	FIRM LN(TOTAL ASSETS)	7.65	1.72	1.48	7.59	15.06
	FIRM CAPITAL RATIO	2.57	1.16	-5.12	2.75	4.61
	FIRM LIQUIDITY RATIO	6.71	11.09	0	2.62	100
	FIRM SUBPRIME	0.10	0.31	0	0	1
	FIRM LN(1+AGE)	1.96	0.92	0	2.08	4.82
	GDP GROWTH	3.54	0.43	2.57	3.64	4.11
	CHANGE IN 3-MONTH INTEREST RATE	0.24	0.82	-1.40	0.11	1.31

ODI GROWIII	5.54	0.43	2.57	5.04	7.11
CHANGE IN 3-MONTH INTEREST RATE	0.24	0.82	-1.40	0.11	1.31
INFLATION	3.12	0.62	2.13	3.15	4.19
Crisis Times (2007:08-2010:06)	Number of Obs	ervations = 98,2	40		
LOAN APPLICATION IS GRANTED	0.30	0.46	0	0	1
BANK LN(TOTAL ASSETS)	17.68	1.42	10.31	17.84	19.94
BANK CAPITAL RATIO	5.33	1.81	0.00	4.85	62.44
BANK LIQUIDITY RATIO	12.54	6.09	0.53	11.74	87.64
BANK DOUBTFUL LOANS RATIO	2.80	2.27	0	2.41	21.93
BANK HERFINDAHL BY INDUSTRY	28.55	8.26	13.65	27.00	100
LN(1+NUMBER OF MONTHS WITH THE BANK)	0.63	1.42	0	0	5.67
FIRM LN(TOTAL ASSETS)	7.74	1.63	1.45	7.70	15.82
FIRM CAPITAL RATIO	2.84	1.04	-4.26	3.01	4.60
FIRM LIQUIDITY RATIO	7.09	10.92	0	3.02	100
FIRM SUBPRIME	0.12	0.33	0	0	1
FIRM LN(1+AGE)	2.26	0.88	0	2.40	4.92
GDP GROWTH	-0.42	2.70	-4.45	-0.86	3.58
CHANGE IN 3-MONTH INTEREST RATE	-0.95	1.85	-4.38	-0.60	1.41
INFLATION	2.14	2.05	-1.37	1.77	5.27

Table 5

The bank and firm balance-sheet channels during good times (2002:02-2007:07) using the loan applications that are also matched with firm balance sheet data

	(1)	(2)	(3)	(4)	(5)
BANK LN(TOTAL ASSETS)	-4.45 ***	-2.12	-1.31		
	(1.60)	(2.19)	(2.58)		
BANK CAPITAL RATIO	-0.21	-0.24	-0.20		
	(0.27)	(0.28)	(0.35)		
BANK LIQUIDITY RATIO	-0.09	-0.14	-0.18 *		
	(0.14)	(0.14)	(0.10)		
BANK DOUBTFUL LOANS RATIO	0.13	0.32	0.82		
	(0.81)	(0.85)	(0.85)		
BANK HERFINDAHL BY INDUSTRY	-0.07	-0.07	-0.10		
	(0.08)	(0.09)	(0.09)		
LN(1+NUMBER OF MONTHS WITH THE BANK)	1.93 ***	1.95 ***	0.80 ***	0.78 ***	0.78 ***
	(0.17)	(0.17)	(0.20)	(0.23)	(0.20)
FIRM LN(TOTAL ASSETS)	-1.32 ***	-1.33 ***	-1.68 *	-1.69	-1.69
	(0.31)	(0.31)	(0.95)	(1.14)	(1.09)
FIRM CAPITAL RATIO	-0.96 ***	-0.94 ***	2.65 ***	2.56 ***	2.56 ***
	(0.22)	(0.22)	(0.62)	(0.87)	(0.71)
FIRM LIQUIDITY RATIO	-0.01	-0.01	0.02	0.03	0.03
	(0.02)	(0.02)	(0.06)	(0.08)	(0.05)
FIRM SUBPRIME	-0.47	-0.41	-7.28 *	-6.93 *	-6.93 *
	(0.59)	(0.59)	(3.75)	(4.01)	(4.08)
FIRM LN(1+AGE)	-1.10 ***	-1.09 ***	-3.67 *	-2.64	-2.64
	(0.35)	(0.36)	(2.22)	(3.02)	(2.44)
GDP GROWTH	3.46 **				
	(1.35)				
CHANGE IN 3-MONTH INTEREST RATE	-0.35				
	(0.68)				
INFLATION	0.02				
	(0.31)				
Fixed Effects					
Bank	Yes	Yes	Yes	-	-
Time	No	Yes	Yes	-	-
Firm	No	No	Yes	Yes	Yes
Bank*Time	No	No	No	Yes	Yes
Number of Observations	100,110	100,110	100,110	100,110	100,110

Notes: The table reports the estimated coefficients and robust standard errors (S.E.) in percent clustered at the bank and firm level in models 1-3, at the firm level in model 4, and at the bank level in model 5 from linear probability models estimated using least squares. The dependent variable is LOAN APPLICATION IS GRANTED. Fixed effects are included ("Yes"), not included ("No"), or comprised by the included set of fixed effects ("-"). The set of time fixed effects includes a fixed effect for every (but one) year:month during the sample period. The variable definitions and summary statistics are in Table 1. ***, ***, and * indicates statistical significant at the 1, 5, and 10 percent level, respectively.

https://www.aeaweb.org/articles?id=10.1257/aer.102.5.2301

Table 6

The bank and firm balance-sheet channels during good (2002:02-2007:07) and crisis times (2007:08-2010:06) using the loan applications that are also matched with firm balance sheet data

firm balance sheet data	(1)	(2)	(3)	(4)	(5)
BANK LN(TOTAL ASSETS)	-0.10	0.02	-0.26	(4)	(3)
DITINE EX (TOTAL PRODUCT)	(2.15)	(3.58)	(3.02)		
BANK CAPITAL RATIO	-0.42	-0.55	-0.52		
BINK CHITTE KITTO	(0.42)	(0.42)	(0.40)		
BANK LIQUIDITY RATIO	-0.05	-0.06	-0.07		
BANK EIQUIDITT KATIO	(0.11)	(0.11)	(0.08)		
BANK DOUBTFUL LOANS RATIO	0.33	0.42	0.38		
BANK DOUBTFUL LOANS KATIO	(0.69)	(0.70)	(0.62)		
BANK HERFINDAHL BY INDUSTRY	-0.05	-0.08	-0.03		
DANK HERPINDAHE DI INDUSTRI	(0.09)	(0.10)	(0.08)		
LN(1+NUMBER OF MONTHS WITH THE BANK)	1.92 ***	1.94 ***	0.84 ***	0.86 ***	0.86 ***
EN(1+NOWIDER OF MONTHS WITH THE BANK)			(0.22)		
FIRM LN(TOTAL ASSETS)	(0.18) -1.35 ***	(0.17) -1.35 ***	-2.65 ***	(0.21) -2.79 ***	(0.22) -2.79 ***
FIRM LN(TOTAL ASSETS)					
FIDM CADITAL DATIO	(0.31)	(0.31)	(0.67)	(0.86)	(0.70)
FIRM CAPITAL RATIO	-0.98 ***	-0.95 ***	1.47 ***	1.51 **	1.51 ***
EIDM I IOUIDITY DATIO	(0.22)	(0.22)	(0.51)	(0.66)	(0.50)
FIRM LIQUIDITY RATIO	-0.01	-0.01	-0.03	-0.03	-0.03
EIDM GUDDDIME	(0.02)	(0.02)	(0.04)	(0.06)	(0.04)
FIRM SUBPRIME	-0.45	-0.38	-4.88 **	-4.56 **	-4.56 **
EIDM LN/1 - A CIEV	(0.58)	(0.58)	(1.95)	(2.27)	(2.05)
FIRM LN(1+AGE)	-1.05 ***	-1.04 ***	-0.94	0.23	0.23
	(0.36)	(0.36)	(1.81)	(2.26)	(1.82)
CRISIS	-37.01 ***				
	(8.81)				
CRISIS * BANK LN(TOTAL ASSETS)	1.15 ***	1.10 ***	1.13 ***		
	(0.38)	(0.38)	(0.34)		
CRISIS * BANK CAPITAL RATIO	1.21 ***	1.20 ***	1.14 ***		
	(0.29)	(0.31)	(0.23)		
CRISIS * BANK LIQUIDITY RATIO	0.17	0.15	0.17 *		
	(0.13)	(0.13)	(0.09)		
CRISIS * BANK DOUBTFUL LOANS RATIO	-0.72	-0.94	-0.91		
	(0.70)	(0.73)	(0.61)		
CRISIS * BANK HERFINDAHL BY INDUSTRY	-0.06	-0.06	-0.10 *		
	(0.09)	(0.08)	(0.06)		
CRISIS * LN(1+NUMBER OF MONTHS WITH THE BANK)	-0.19	-0.21	-0.06	-0.13	-0.13
	(0.20)	(0.19)	(0.25)	(0.27)	(0.25)
CRISIS * FIRM LN(TOTAL ASSETS)	-0.11	-0.11	-0.50	-0.38	-0.38
	(0.32)	(0.32)	(0.41)	(0.41)	(0.45)
CRISIS * FIRM CAPITAL RATIO	0.97 ***	0.95 ***	2.60 ***	2.40 ***	2.40 ***
	(0.33)	(0.32)	(0.49)	(0.60)	(0.49)
CRISIS * FIRM LIQUIDITY RATIO	0.00	0.00	-0.01	-0.01	-0.01
	(0.02)	(0.02)	(0.06)	(0.07)	(0.07)
CRISIS * FIRM SUBPRIME	-0.49	-0.65	0.23	0.08	0.08
	(0.68)	(0.68)	(1.32)	(1.62)	(1.42)
CRISIS * FIRM LN(1+AGE)	2.20 ***	2.21 ***	1.69 **	1.86 **	1.86 **
	(0.37)	(0.37)	(0.71)	(0.85)	(0.79)
GDP GROWTH	1.74 ***	, ,	, ,	, ,	, ,
	(0.48)				
CHANGE IN 3-MONTH INTEREST RATE	-1.04 *				
	(0.56)				
INFLATION	-0.39				
·- 	(0.27)				
Fixed Effects	(*/				
Bank	Yes	Yes	Yes	-	-
Time	No	Yes	Yes	-	-
Firm	No	No	Yes	Yes	Yes
Bank*Time	No	No	No	Yes	Yes
Number of Observations	198,350	198,350	198,350	198,350	198,350

Notes: The table reports the estimated coefficients and robust standard errors (S.E.) in percent clustered at the bank and firm level in models 1-3, at the firm level in model 4, and at the bank level in model 5 from linear probability models estimated using least squares. The dependent variable is LOAN APPLICATION IS GRANTED. Fixed effects are included ("Yes"), not included ("No"), or comprised by the included set of fixed effects ("-"). The set of time fixed effects includes a fixed effect for every (but one) year:month during the sample period. The variable definitions and summary statistics are in Table 1. ***, **, and * indicates statistical significant at the 1, 5, and 10 percent level, respectively.